

RANDOM WALK IN 1 DIMENSION

MODEL: N steps of ± 1 along x -axis,
with $P(+1) = P(-1) = \frac{1}{2}$.

Probability of k steps to right and $N-k$ to left is

$$P_k(N; \frac{1}{2}) = \binom{N}{k} \left\{ \frac{1}{2} \right\}^N \quad \text{BINOMIAL DISTRIBUTION}$$

Displacement [from starting point] is related to k by

$$x = \underset{\substack{\uparrow \\ \text{to right}}}{k} - \underset{\substack{\downarrow \\ \text{to left}}}{(N-k)} = 2k - N$$

Expectation value [obvious by symmetry]:

$$\begin{aligned} \langle x \rangle &= \langle 2k - N \rangle \\ &= 2\langle k \rangle - N = 2 \times \frac{N}{2} - N = 0. \end{aligned}$$

Variance of x :

$$\begin{aligned} \text{var}(x) &= \langle x^2 \rangle = \langle (2k - N)^2 \rangle \\ &= 4 \text{var } k = 4 \times N \times \frac{1}{2} \times \frac{1}{2} \\ &= N. \end{aligned}$$

$$\Rightarrow \sigma_x = x_{\text{RMS}} = \sqrt{N}.$$

PL10471: LECTURE 20

[RANDOM WALK IN 1D]

The probability $P(x; N)$ that the walk has reached x after N steps is [of course] the same thing as $P_k(N; 1/2)$ with $k = (x+N)/2$.

From the fact that P_k is Gaussian for large N it follows that $P(x; N)$ is also Gaussian in this limit:

$$P(x; N) \approx \frac{e^{-x^2/2\sigma_x^2}}{(2\pi\sigma_x^2)^{1/2}} = \frac{e^{-x^2/2N}}{(2\pi N)^{1/2}}$$

Valid for $N \gg 1$ and $x \ll N$.

"GOING NOWHERE" IN 1D

A random walk in 1D is certain to return to its starting point — eventually!

PROOF relies on the symmetry of $P(x; N)$ about $x=0$ and the fact that the probability of remaining in the range $[-L, L]$ tends to zero for $N \rightarrow \infty$:

$$P(-L \leq x \leq L; N)$$

$$= \sum_{x=-L}^L P(x; N)$$

$$\approx \sum_{x=-L}^L \frac{e^{-x^2/2N}}{\sqrt{2\pi N}} \quad \text{GAUSSIAN LIMIT, } N \gg 1.$$

$$< \frac{2L+1}{\sqrt{2\pi N}} \rightarrow 0 \text{ as } N \rightarrow \infty.$$

Hence the particle is certain to exit from any region, and if it started from the middle, the probability of exiting to left or right is $1/2$ [by symmetry].

["GOING NOWHERE" PROOF]

STRATEGY: We show that the prob. of NOT returning to $x=0$ can be broken down to a sequence of independent exit events whose prob. tends to 0.

STEP 1: The particle is at $x=1$ [or -1] after the first rand. step away from $x=0$. On its next step, it either returns to $x=0$ or "exits" to $x=2$.

After STEP 1, the probability that it's at $x=2$ is $1/2$.

STEP 2: 


Particle starts at centre of this region.

The particle has room to move around, but will eventually exit to $x=0$ or $x=4$. The prob. after STEP 2 that it has not yet returned is $\frac{1}{2} \times \frac{1}{2}$

prob. it
hadn't returned
to 0 after Step 1

prob. it exits
to right in STEP 2

["GOING NOWHERE" PROOF]

STEP 3: 

particle starts at
centre of this region.

Eventually it exits, either to $x=0$ or $x=8$. The prob. after STEP 3 that it has not yet returned to $x=0$ is $\frac{1}{2} \times \frac{1}{2} \times \frac{1}{2}$.

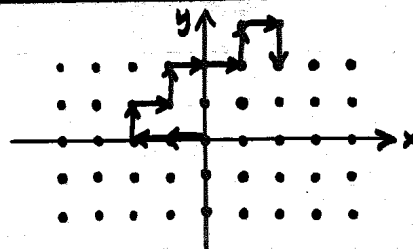
AND SO ON. After step N , the prob. that that the particle has not returned to $x=0$ will be $(\frac{1}{2})^N \rightarrow 0$ for $N \rightarrow \infty$.

Hence the probability of return to the starting point $= 1 - \frac{1}{2^N} \rightarrow 1$.

[Aside: A similar result holds for random walks in 2D.

There is no such result for 3D: the probability of return is non-zero, but less than 1.]

RANDOM WALK IN 2D



Each step is $\uparrow, \downarrow, \rightarrow, \leftarrow$ with equal probability.

We want to find $P(x, y; N)$, the prob. that we are at (x, y) after N steps.

Break this up: prob. that M steps are \uparrow or \downarrow and $N-M$ are \leftarrow or \rightarrow is $P_M(N; \frac{1}{2})$.

Given M steps \uparrow, \downarrow , the probability of reaching y is $P(y; M)$, since each of the steps is equally likely to be \uparrow or \downarrow [so it's a random walk of M steps in y].

Similarly, the prob. of reaching x in $N-M$ steps is $P(x; N-M)$.

[2D RANDOM WALK, CONT'D]

Hence

$$P(x, y; N) = \sum_{M=0}^N P(x; N-M) P(y; M) P_M(N, \frac{1}{2})$$

[We sum over the mutually exclusive possibilities for M.]

Looks messy, but simplifies for $N \gg 1$, because $P_M(N, \frac{1}{2})$ is sharply peaked around $M = N/2$, with width $\propto \sqrt{N} \ll N$.

$P(x; N-M) P(y; M)$ varies very little over this range $\langle M \rangle \pm \sqrt{N}$, so can treat that as a [nearly!] constant factor in the sum:

$$P(x, y; N) \approx P(x; N - \underbrace{\langle M \rangle}_{\frac{N}{2}}) P(y; \underbrace{\langle M \rangle}_{\frac{N}{2}}) \underbrace{\sum_M P_M(N; \frac{1}{2})}_1$$

$$= P(x; \frac{N}{2}) P(y; \frac{N}{2})$$

$$\propto \frac{e^{-x^2/N}}{\sqrt{\pi N}} \cdot \frac{e^{-y^2/N}}{\sqrt{\pi N}}$$

$$= \frac{e^{-(x^2+y^2)/N}}{\pi N} \quad \text{DEPENDS ONLY ON DIST. FROM ORIGIN.}$$

EXTENSION TO 3D

Can progress from 2D to 3D in much the same way as we got from 1D to 2D.

The probability that any given step is in the $\pm z$ direction is $1/3$, so prob. that M out of N steps are $\pm z$ is $P_M(N; 1/3)$, which is sharply peaked around $M = N/3$.

Given M steps $\pm z$, the probability of reaching x, y, z is $P(x, y; N-M) P(z; M)$.

Hence

$$P(x, y, z; N) = \sum_{M=0}^N P(x, y; N-M) P(z; M) P_M(N; 1/3) \\ = P(x, y; N - \frac{N}{3}) P(z; \frac{N}{3}) \underbrace{\sum_M P_M}_{1}$$

$$= P(x, y; \frac{2}{3}N) P(z; \frac{N}{3})$$

$$= P(x; \frac{N}{3}) P(y; \frac{N}{3}) P(z; \frac{N}{3})$$

$$= \frac{e^{-(x^2+y^2+z^2)/(2N/3)}}{(2\pi N/3)^{3/2}}$$